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A note on an inequality of Chernoff^{*)}

by

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ABSTRACT

An inequality due to Chernoff is generalized.

KEY WORDS & PHRASES: *Inequality of Chernoff, Cauchy-Schwarz inequality*

^{*)} This report will be submitted for publication elsewhere.

1. INTRODUCTION

Let X be a standard normal random variable and let $G : \mathbb{R} \rightarrow \mathbb{R}$ be a function which is absolutely continuous with respect to Lebesgue measure with Radon-Nikodym derivative g . In CHERNOFF (1980,1981) the elegant inequality

$$(1.1) \quad \text{var } G(X) \leq E g^2(X)$$

has been presented. For an arbitrary random variable X we shall derive the generalization (2.7) of (1.1). Although this derivation has not been restricted to the normal case, it seems to be somewhat simpler than the one given in CHERNOFF (1981). The main idea in the proof of (2.7), namely the use of the Cauchy-Schwarz inequality, is also contained in CACOULLOS (1981), which restricts attention to the case $h = 1$, $\alpha = c = 0$ in Theorem 2.1, and in CHEN (1980), which considers the multivariate normal case.

2. THE RESULT AND SOME EXAMPLES

For our formulation of Chernoff's inequality we need the convention that the variance of a random variable is infinite iff the second moment of that random variable is infinite. Furthermore, for $a \leq b$ we'll denote an integral over $(a, b]$ by \int_a^b or $-\int_b^a$.

THEOREM 2.1. *Let μ be a σ -finite measure on $(\mathbb{R}, \mathcal{B})$ and let X be a random variable with density f with respect to μ . S is the smallest interval containing the support of f and $h : S \rightarrow \mathbb{R}$ is a nonnegative measurable function such that for all $a, b \in S$*

$$(2.1) \quad \left| \int_a^b h d\mu \right| < \infty$$

and

$$(2.2) \quad E_f \left| \int_a^X h d\mu \right| < \infty.$$

Now there exist $c \in S$ and $\alpha \in [0,1]$ such that

$$(2.3) \quad E_f H(X) = 0$$

holds for

$$(2.4) \quad H(x) = \int_c^x h d\mu + \alpha h(c) \mu(\{c\}), \quad x \in S.$$

Let $g : S \rightarrow \mathbb{R}$ be a measurable function satisfying (2.1) and for some $d \in S$ and $e \in \mathbb{R}$ define

$$(2.5) \quad G(x) = \int_d^x g d\mu + e, \quad x \in S.$$

If

$$(2.6) \quad \mu(\{x \in S \mid g(x) \neq 0, f(x)h(x) = 0\}) = 0,$$

then the inequality

$$(2.7) \quad \text{var}_f G(X) \leq E_f \left\{ \frac{g^2(x)}{f(x)h(x)} \int_{S \cap [X, \infty)} H f d\mu \right\}$$

is valid, with equality iff the variance is infinite or g is a multiple of h μ -almost everywhere on S .

The proof of this theorem will be presented in section 3. In Table 2.1 the upperbound B from (2.7) is given for some choices of f and H with μ Lebesgue measure. Note that (1.1) can be obtained from (2.7) by choosing $h(x) = 1$, $x \in \mathbb{R}$ (cf. examples 1 and 6 of Table 2.1).

name	f	H	B
1 normal	$(2\pi\sigma^2)^{-\frac{1}{2}} e^{-\frac{1}{2}x^2\sigma^{-2}}$	x	$\sigma^2 E_f g^2(X)$
2 exponential	$\lambda e^{-\lambda x}, x > 0$	$(x - \frac{2}{\lambda}) e^{\frac{1}{2}\lambda x}$	$4\lambda^{-2} E_f g^2(X)$
3 Laplace	$\frac{1}{2}\lambda e^{-\lambda x }$	$x e^{\frac{1}{2}\lambda x }$	$4\lambda^{-2} E_f g^2(X)$
4 logistic	$2\lambda(e^{\lambda x} + e^{-\lambda x})^{-2}$	$e^{\lambda x} - e^{-\lambda x}$	$\lambda^{-2} E_f g^2(X)$
5 gamma	$[\sigma^\alpha \Gamma(\alpha)]^{-1} x^{\alpha-1} e^{-\frac{x}{\sigma}}, x > 0$	$x - \alpha\sigma$	$\sigma E_f X g^2(X)$
6 -	$c(\alpha, \sigma) x ^{\alpha-1} e^{-\frac{1}{2} \frac{x}{\sigma} ^{\alpha+1}}$	x	$\frac{2\sigma^{\alpha+1}}{\alpha+1} E_f X ^{1-\alpha} g^2(X)$

Table 2.1. The value of the upperbound B of (2.7) for some choices of f and H with μ Lebesgue measure.

For μ counting measure on the integers a few examples of (2.7) are given in Table 2.2. We only note the following. If X has a discrete distribution with mean ν then $H(x) = x - \nu, x \in \mathbb{Z}$, is realized by the choices $h(x) = 1, c = [\nu] + 1$ and $\alpha = [\nu] + 1 - \nu$, where $[\nu]$ denotes the integer part of ν (cf. (2.3) and (2.4)).

name	f	H	B
1 Poisson	$e^{-\lambda} \lambda^x (x!)^{-1}$	$x - \lambda$	$E_f X g^2(X)$
2 binomial	$\binom{n}{x} p^x (1-p)^{n-x}$	$x - np$	$(1-p) E_f X g^2(X)$
3 negative binomial	$\binom{x-1}{k-1} p^k (1-p)^{x-k}$	$x - \frac{k}{p}$	$p^{-1} E_f (X-k) g^2(X)$

Table 2.2. The value of the upperbound B of (2.7) for some choices of f and H with μ counting measure on the integers.

3. PROOF

If h vanishes μ -almost everywhere on S then (2.3) and (2.4) are fulfilled for all $c \in S$ and $\alpha \in [0,1]$. Therefore, for the proof of (2.3) we assume without loss of generality that h does not have this property. In view of (2.1) and (2.2) we can define $\psi : S \rightarrow \mathbb{R}$ by

$$\psi(a) = E_f \int_a^X h d\mu.$$

Since h is nonnegative ψ is nonincreasing on S . Because $\psi(a)$ is finite, Fubini's theorem yields

$$(3.1) \quad \psi(a) = - \int_{S \cap (-\infty, a]} \int_{S \cap (-\infty, y)} f(x) d\mu(x) h(y) d\mu(y) \\ + \int_{S \cap (a, \infty)} \int_{S \cap [y, \infty)} f(x) d\mu(x) h(y) d\mu(y).$$

Let $a_0 = \inf S$. If $a_0 \in S$ we obtain the nonnegativity of $\psi(a_0)$ from (3.1). If $a_0 \notin S$, then we see by the dominated and the monotone convergence theorem that

$$\lim_{a \downarrow a_0} \psi(a) = \int_S \int_{S \cap [y, \infty)} f(x) d\mu(x) h(y) d\mu(y) > 0.$$

In both cases there exists a $c_0 \in S$ with $\psi(c_0) \geq 0$. Analogously we see that there exists a $c_1 \in S$ with $\psi(c_1) \leq 0$. If there is a $c \in [c_0, c_1]$ with $\psi(c) = 0$, then (2.3) and (2.4) are valid with $\alpha = 0$. If not, there exists a $c \in (c_0, c_1]$ such that $0 \leq \psi(c-) < \infty$ and $-\infty < \psi(c) < 0$ and hence also an $\alpha \in (0,1]$ such that

$$(3.2) \quad \alpha \psi(c-) + (1-\alpha)\psi(c) = 0.$$

In view of (2.4) equality (3.2) is the same one as (2.3).

By the Cauchy-Schwarz inequality, Fubini's theorem and (2.3) we obtain for g satisfying (2.5) and (2.6)

$$\begin{aligned}
\text{var}_f G(X) &\leq E_f \left[\int_c^X \frac{g}{h^{\frac{1}{2}}} h^{\frac{1}{2}} d\mu + \alpha \frac{g(c)}{h^{\frac{1}{2}}(c)} h^{\frac{1}{2}}(c) \mu(\{c\}) \right]^2 \\
&\leq E_f \left[\int_c^X \frac{g^2}{h} d\mu + \alpha \frac{g^2(c)}{h(c)} \mu(\{c\}) \right] H(X) \\
&= \int_S \int_c^x \frac{g^2(y)}{h(y)} d\mu(y) H(x) f(x) d\mu(x) \\
&= - \int_{S_n(-\infty, c]} \int_{S_n(-\infty, y)} H(x) f(x) d\mu(x) \frac{g^2(y)}{h(y)} d\mu(y) \\
&\quad + \int_{S_n(c, \infty)} \int_{S_n[y, \infty)} H(x) f(x) d\mu(x) \frac{g^2(y)}{h(y)} d\mu(y) \\
&= \int_S \frac{g^2(y)}{f(y)h(y)} \int_{S_n[y, \infty)} H(x) f(x) d\mu(x) f(y) d\mu(y),
\end{aligned}$$

which implies (2.7) and thereby the theorem. \square

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REFERENCES

- CACOULLOS, T. (1981), *On upper and lower bounds for the variance of a function of a random variable*, to appear.
- CHEN, L.H.Y. (1980), *An inequality for the multivariate normal distribution*, M.I.T. Technical Report No. 18.
- CHERNOFF, H. (1980), *The identification of an element of a large population in the presence of noise*, *Ann. Statist.* 8, 1179-1197.
- CHERNOFF, H. (1981), *A note on an inequality involving the normal distribution*, *Ann. Probab.* 9, 533-535.

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