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On the attained waiting time

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## On the Attained Waiting Time

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### Summary

By using properties of up- and downcrossings of the sample functions of the work load process and of the attained waiting time process for a G/G/1 queueing model it is shown that both processes have the same stationary distribution, if such distributions do exist.

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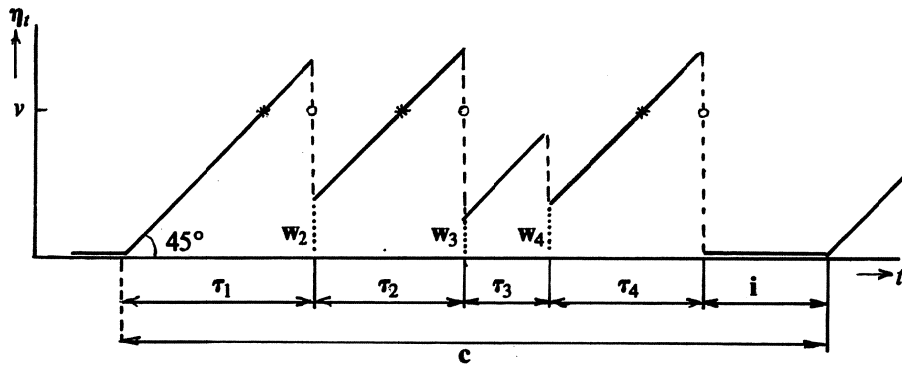
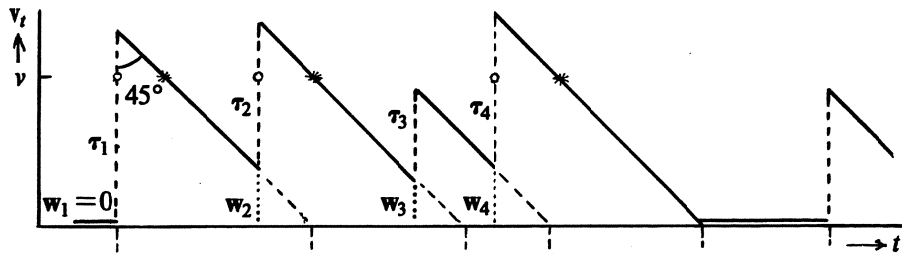
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Sakasegawa and Wolff [1] show by using sample function arguments that for the FIFO G/G/1 queueing model the workload process  $v_t$  and the attained waiting time process  $\eta_t$  possess the same stationary distribution, if such distributions exist. However their proof is somewhat artificial (see their use of preemptive LIFO).

A direct proof proceeds as follows. Consider a busy cycle  $c$  with  $n$  the number of customers served;  $\tau_1, \dots, \tau_n$  are the service times of these customers,  $w_1, \dots, w_n$  their successive actual waiting times,  $i$  is the idle time, so

$$c = \tau_1 + \dots + \tau_n + i \quad (1)$$

The attained service time  $\eta_t$  at epoch  $t$  is by definition the time between  $t$  and the arrival epoch of the customer being served at epoch  $t$ . In the figure below the sample function of the work load process  $v_t$  and the corresponding  $\eta_t$ -process during the busy cycle  $c$  are shown, with  $n=4$ .



Define for  $v \geq 0$ ,

$$d(v) := \# \text{ downcrossings of } v_t, 0 \leq t \leq c \text{ with level } v, (*) \quad (2)$$

$$u(v) := \# \text{ upcrossings } ,, v_t, 0 \leq t \leq c ,, ,, v, (o) \quad (2)$$

$$\delta(v) := \# \text{ upcrossings of } \eta_t, 0 \leq t \leq c ,, ,, v, (*) \quad (3)$$

$$\omega(v) := \# \text{ downcrossings } ,, \eta_t, 0 \leq t \leq c ,, ,, v, (o). \quad (3)$$

Note that in the figure  $d(v)=3$ ; the upcrossings are there indicated by  $o$ , the downcrossings by  $*$ . It is immediately evident from the geometry of the sample functions, cf. [2], [3], that with probability one, for  $v \geq 0$ ,

$$d(v) = u(v), \quad \delta(v) = \omega(v), \quad (4)$$

$$u(v) = \delta(v); \quad (5)$$

and

$$d(v) = \frac{d}{dv} \int_0^c (v_t < v) dt, \quad \delta(v) = \frac{d}{dv} \int_0^c (\eta_t < v) dt, \quad (6)$$

where we use the notation

$$(v_t < v) \equiv 1_{v_t < v} \text{ and } \int_0^c (v_t < v) dt \equiv \int_0^{\infty} (v_t < v, c \geq t) dt, \quad (7)$$

for the indicator function and the integral. Since

$$i = \left\{ \int_0^c (v_t < v) dt \right\}_{v=0+} = \left\{ \int_0^c (\eta_t < v) dt \right\}_{v=0+}, \quad (8)$$

integration of (6), using the boundary conditions (8) yields, via (4) and (5), that with prob. 1,

$$\int_0^c (v_t < v) dt = \int_0^c (\eta_t < v) dt, \quad v \geq 0. \quad (9)$$

Because

$$(v_t < v) = 1 - (v_t \geq v),$$

we have from (9)

$$\int_0^c (v_t \geq v) dt = \int_0^c (\eta_t \geq v) dt,$$

which is theorem 1 of [1].

For the GI/G/1 queueing model with the conditions:

i.  $E\{c\} < \infty$ ,

ii.  $c$  has not a lattice distribution;

the stochastic mean value theorem, cf. [3], [4], applies, i.e. the  $v_t$ -process has a unique stationary distribution and for  $v_\infty$  a stochastic variable having this distribution holds

$$\Pr\{v_\infty < v\} = \frac{1}{E\{c\}} E\left\{ \int_0^c (v_t < v) dt \right\}, \quad v \geq 0, \quad (10)$$

For the same conditions it is similarly shown that the  $\eta_t$ -process possesses a stationary distribution and for  $\eta_\infty$  a stochastic variable with this distribution holds

$$\Pr\{\eta_\infty < v\} = \frac{1}{E\{c\}} E\left\{ \int_0^c (\eta_t < v) dt \right\}. \quad (11)$$

Consequently from (10) and (11),

$$\eta_\infty \sim v_\infty,$$

a result obtained in [5]. By using again the properties of up- and downcrossings it is readily shown that for the G/G/1 queue the limits for  $T \rightarrow \infty$  of

$$\frac{1}{T} \int_0^T (v_t < v) dt \text{ and } \frac{1}{T} \int_0^T (\eta_t < v) dt, \quad v \geq 0,$$

both exist with probability one and are equal with probability one (note that the number of upcrossings and that of downcrossings in an interval  $(0, T)$  differ by at most one).

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