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# Heavy-Traffic Theory for the Heavy-Tailed M/G/1 Queue and $\nu$ -stable Lévy Noise Traffic

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## ABSTRACT

The workload  $\mathbf{v}_t$  of an M/G/1 model with traffic  $a < 1$  is analyzed for the case with heavy-tailed message length distributions  $B(\tau)$ , e.g.  $1 - B(\tau) = O(\tau^{-\nu})$ ,  $\tau \rightarrow \infty$ ,  $1 < \nu \leq 2$ . It is shown that a factor  $\Delta(a)$  exists with  $\Delta(a) \downarrow 0$  for  $a \uparrow 1$  such that, whenever  $\mathbf{v}_t$  is scaled by  $\Delta(a)$  and time  $t$  by  $\Delta_1(a) = \Delta(a)(1 - a)$  then  $\mathbf{w}_\tau(a) = \Delta(a)\mathbf{v}_{\tau/\Delta_1(a)}$  converges in distribution for  $a \uparrow 1$  and every  $\tau > 0$ . Proper scaling of the traffic load  $\mathbf{k}_t$ , generated by the arrivals in  $[0, t)$ , leads to

$$\tilde{\mathbf{w}}_\tau = \max[\mathbf{H}(\tau), \sup_{0 < u < \tau} (\mathbf{H}(\tau) - \mathbf{H}(u))], \quad \tau > 0,$$

with  $\mathbf{H}(\tau) = \mathbf{N}(\tau) - \tau$ . Here  $\{\mathbf{N}(\tau), \tau \geq 0\}$  with  $\nu \neq 2$  is  $\nu$ -stable Lévy motion, for  $\nu = 2$  it is Brownian motions and  $\tilde{\mathbf{w}}_\tau$  has the limiting distribution of  $\mathbf{w}_\tau(a)$  for  $a \uparrow 1$ . This relation is analogous to Reich's formula for the M/G/1 model with  $a < 1$ . The results obtained are generalisations of the diffusion approximation of the M/G/1 model with  $B(\tau)$  having a finite second moment.

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## 1. INTRODUCTION

In this study we analyse the workload process  $\{\mathbf{v}_t, t \geq 0\}$  of an M/G/1 model with traffic load  $a$  and with heavy-tailed message length distribution  $B(t)$ . For instance

$$1 - B(t) = O(t^{-\nu}), \quad t \rightarrow \infty, \quad 1 < \nu \leq 2,$$

is an example of the class of message length distributions to be considered, see also [6] for examples. For such an M/G/1 model explicit representations have recently been obtained for the LST of the stationary waiting time distribution; in particular for  $\nu = 1\frac{1}{2}$  also for the distribution function, see [1], [7]. In [1] it has been shown that whenever  $\mathbf{v}$  is a stochastic variable with distribution the stationary waiting time distribution ( $a < 1$ ) then a unique scaling factor  $\Delta(a)$ , the contraction factor, exists such that  $\Delta(a)\mathbf{v}$  converges in distribution for  $a \uparrow 1$  to a true probability distribution. Actually this formulates a heavy traffic theorem for  $a \uparrow 1$  for the M/G/1 model with heavy-tailed message length distributions. This result has been extended for the GI/G1 queue, see [8].

In the present study we analyse the workload process  $\mathbf{v}_t$  of the heavy-tailed M/G/1 queueing model for  $a \uparrow 1$ . It will be shown that when the workload  $\mathbf{v}_t$  is scaled by  $\Delta(a)$  and the time  $t$  by  $\Delta_1(a) = \Delta(a)(1 - a)$  then  $\mathbf{w}_\tau(a) := \Delta(a)\mathbf{v}_{\tau/\Delta_1(a)}$  converges in distribution for  $a \uparrow 1$  for every  $\tau > 0$ . It is further shown that the so scaled or contracted workload process  $\{\mathbf{w}_\tau(a), \tau \geq 0\}$  converges weakly to a workload process of a "queueing model" of which the input process is described by a stable Lévy motion for  $1 < \nu < 2$  and by a Brownian motion for  $\nu = 2$ .

We continue this introduction with an overview of the various sections. In Section 2 we describe the class of heavy-tailed message length distributions by a characterisation of their LST  $\beta(\rho)$ . The essential feature is that  $\beta(\rho)$  varies regularly at  $\rho = 0$  with index  $\nu$ ,  $1 < \nu \leq 2$ . In this section we also

consider the busy period  $\mathbf{p}(\mathbf{v})$  with initial workload  $\mathbf{v}$ , in particular when  $\mathbf{v}$  is a random variable with distribution that of the stationary workload process. It leads to a new functional equation, which is used in Section 3 to study the contracted busy period  $\Delta_1(a)\mathbf{p}(\mathbf{v})$  for  $a \uparrow 1$ . Further Section 2 quotes some known relations concerning the workload  $\mathbf{v}_t$ .

In Section 3 the contracted variables  $\Delta(a)\mathbf{v}$ ,  $\Delta_1(a)\mathbf{p}_0$  and  $\Delta_1(a)\mathbf{p}(\mathbf{v}/\Delta(a))$  are defined, here  $\mathbf{v}$  has the stationary waiting time distribution. It is shown that they converge in distribution for  $a \uparrow 1$ . Their limiting distributions are described.

In Section 4 the variable  $\mathbf{w}_\tau(a) = \Delta(a)\mathbf{v}_\tau/\Delta_1(a)$  is analysed by starting from the relations for the  $\mathbf{v}_t$ -process. It turns out that the scaling of the workload by  $\Delta(a)$  and the scaling of time by  $\Delta_1(a)$ , is the appropriate one. It is shown that  $\mathbf{w}_\tau(a)$  converges in distribution for  $a \uparrow 1$  for every  $\tau > 0$ . For  $\mathbf{w}_\tau$  a stochastic variable with distribution the limiting distribution of  $\mathbf{w}_\tau(a)$  relations are derived which are analogous to those for  $\mathbf{v}_t$  mentioned in Section 2. Theorem 4.1 describes these results.

In Section 5 we analyse the contracted version  $\Delta(a)\mathbf{k}_{\tau/\Delta_1(a)}$  of the input process  $\{\mathbf{k}_t, t > 0\}$ , with  $\mathbf{k}_t$  the amount of traffic generated by the arrivals in  $[0, t)$ . We introduce here the virtual backlog  $\mathbf{h}_t = \mathbf{k}_t - t$  and the noise-traffic  $\mathbf{n}_t = \mathbf{k}_t - at$ . The scaled versions of  $\mathbf{h}_t$  and  $\mathbf{n}_t$  are

$$\mathbf{H}(\tau; a) := \Delta(a)\mathbf{h}_{\tau/\Delta_1(a)}; \mathbf{N}(\tau; a) := \Delta(a)\mathbf{n}_{\tau/\Delta_1(a)},$$

and it shown that

$$\mathbf{H}(\tau; a) = \mathbf{N}(\tau, a) - \tau.$$

$$\mathbf{w}_\tau(a) = \max[\mathbf{H}(\tau; a), \sup_{0 < u < \tau} (\mathbf{H}(\tau; a) - \mathbf{H}(u; a))], \mathbf{w}_0(a) = 0.$$

It is shown that  $\mathbf{H}(\tau; a)$  and  $\mathbf{N}(\tau, a)$  converge in distribution for  $a \uparrow 1$ , their limiting distributions are stable distributions;  $\mathbf{N}(\tau)$  is a variable with distribution the limiting one of  $\mathbf{N}(\tau; a)$ , analogously for  $\mathbf{H}(\tau)$ . The stochastic process  $\{\mathbf{N}(\tau), t \geq 0\}$  appears to be a  $\nu$ -stable Lévy motion for  $1 < \nu < 2$ , and a Brownian motion for  $\nu = 2$ , further,

$$\mathbf{H}(\tau) = \mathbf{N}(\tau) - \tau, \text{ a.s.}$$

With  $\tilde{\mathbf{w}}_\tau$  defined by

$$\tilde{\mathbf{w}}_\tau := \max[\mathbf{H}(\tau), \sup_{0 < u < \tau} (\mathbf{H}(\tau) - \mathbf{H}(u))],$$

it is shown that the  $\{\mathbf{w}_\tau(a), \tau \geq 0\}$  process converges weakly for  $a \uparrow 1$  to the  $\{\tilde{\mathbf{w}}_\tau, \tau \geq 0\}$  process.

For the M/G/1-model the workload process  $\mathbf{v}_t$  is described by Reich's formula

$$\mathbf{v}_t = \max[\mathbf{h}_t, \sup_{0 < u < t} (\mathbf{h}_t - \mathbf{h}_u)], \mathbf{v}_0 = 0.$$

For the contracted M/G/1 model, i.e. work scaled by  $\Delta(a)$  and time by  $\Delta_1(a)$ , the relation for  $\mathbf{v}_t$  transforms into that of  $\mathbf{w}_\tau(a)$  and here the noise traffic  $\mathbf{N}(\tau; a)$  is seen to be the equivalent to the noise traffic  $\mathbf{n}_t$  for the unscaled M/G/1 model. This holds also in the limit for  $a \uparrow 1$ , and for this limiting case the  $\nu$ -stable Lévy motion  $\{\mathbf{N}(\tau), \tau \geq \nu\}, 1 < \nu \leq 2$ , represents the noise traffic of the limiting M/G/1 model.

In appendix A an asymptotic expression for the tail of the limiting distribution of  $\Delta_1(a)\mathbf{p}_0$  for  $a \uparrow 1$  is derived.

The results obtained in the present study are extensions of the heavy traffic theory for the M/G/1 model with  $B(t)$  having a finite second moment. For such an M/G/1 model it is known, that the diffusion approximation provides good results from the numerical viewpoint, even for traffic  $a$  not so close to one. For the heavy-tailed M/G/1 model the equivalent approximation is here the process with traffic input the noise-traffic  $\mathbf{N}(\tau)$ , the  $\nu$ -stable Lévy motion. It is conjectured that such an approach will yield a good approximation. This is motivated by the numerical results in [9] for the waiting time distribution when it is approximated by the distribution of the contracted waiting time.

## 2. THE MODEL

We consider the M/G/1-queue with arrival rate  $\lambda$  and message length distribution  $B(t)$  with

$$\beta := \int_0^{\infty} \tau dB(\tau) < \infty.$$

and

$$a := \lambda\beta < 1. \quad (2.1)$$

In our analysis  $\beta$  will be taken as the unit of time. We consider message length distributions of the following type. For a finite  $T > 0$  it is assumed that: for  $t > T$ ,

$$1 - B(t) = \mathcal{G}_1(t) + \mathcal{G}_2(t). \quad (2.2)$$

Here  $\mathcal{G}_2(t)$  is a function with the property

$$\left| \int_T^{\infty} e^{-\rho t} d\mathcal{G}_2(t) \right| < \infty \text{ for } \operatorname{Re} \rho > -\delta, \text{ for a } \delta > 0, \quad (2.3)$$

and the function  $\mathcal{G}_1(t)$  is the dominant part of the tail of  $1 - B(t)$  for  $t \rightarrow \infty$  with, (2.2);

$$\mathcal{G}_1(t) = O(t^{-\nu}), \quad t \rightarrow \infty, \quad 1 < \nu \leq 2, \quad (2.4)$$

is an example of a heavy-tailed message length distribution.

Denote by  $\beta(\rho)$  the Laplace-Stieltjes transform (LST) of  $B(t)$ ,

$$\beta(\rho) := \int_{0-}^{\infty} e^{-\rho t} dB(t), \quad \operatorname{Re} \rho \geq 0. \quad (2.5)$$

For distributions of the type (2.2) the LST  $\beta(\rho)$  can be described by

$$1 - \frac{1 - \beta(\rho)}{\rho} = g(\rho) + c\rho^{\nu-1}L(\rho) \text{ for } \operatorname{Re} \rho \geq 0, \quad (2.6)$$

with

- i*  $c > 0$  a constant
  - ii*  $1 < \nu \leq 2$ ,
  - iii*  $g(\rho)$  is a regular function of  $\rho$  for  $\operatorname{Re} \rho > -\delta$ ,
  - iv*  $L(\rho)$  is regular for  $\operatorname{Re} \rho > 0$ , and continuous for  $\operatorname{Re} \rho \geq 0$ , except possibly at  $\rho = 0$ ,  
 $L(\rho) \rightarrow b > 0$  for  $|\rho| \rightarrow 0$ ,  $\operatorname{Re} \rho \geq 0$  with  $b = \infty$  if  $\nu = 2$ ,
- (2.7)

$$\lim_{x \downarrow 0} \frac{L(\rho x)}{L(x)} = 1 \text{ for } \operatorname{Re} \rho \geq 0, \quad \rho \neq 0,$$

REMARK 2.1. The principal value of  $\rho^{\nu-1}$  is defined so that  $\rho^{\nu-1} > 0$  for  $\rho > 0$ . □

For examples of distributions  $B(t)$  of the type (2.3) and with LST as given by (2.6), see [1]; cf. further [2], vol. I, p. 467, 501.

Note, that (2.6) and (2.7) imply, that: for  $\operatorname{Re} \rho \geq 0$ ,

$$g(\rho) \rightarrow 0 \text{ and } \rho^{\nu-1}L(\rho) \rightarrow 0 \text{ for } |\rho| \rightarrow 0. \quad (2.8)$$

The contraction coefficient  $\Delta(a)$  is defined as that zero of

$$x^{\nu-1}L(x) = \frac{1-a}{ac}, \quad x > 0, \quad (2.9)$$

which tends to zero for  $a \uparrow 1$ . From (2.6),..., (2.9), it is not difficult to see that  $\Delta(a)$  is well defined and that it is a simple zero. Note, herefore, that the lefthand side of (2.6) is an increasing function of  $\rho \in [0, \infty)$ , and that  $\Delta(a)$  is real for  $(1-a)/ac < 1$  or for  $0 < 1-a \ll 1$ .

Our present analysis mainly concerns the workload  $\mathbf{v}_t$  at time  $t > 0$ . It is well known that for  $a < 1$  this stochastic variable  $\mathbf{v}_t$  converges in distribution for  $t \rightarrow \infty$ . The limiting distribution is the stationary waiting time distribution. Let  $\mathbf{v}$  be a stochastic variable with distribution this stationary distribution, then, as is well known, we have for  $\text{Re } \rho \geq 0$ ,

$$\mathbb{E}\{e^{-\rho \mathbf{V}}\} = \frac{1-a}{1-a\frac{1-\beta(\rho)}{\rho}} = \frac{1}{1+\frac{a}{1-a}\{1-\frac{1-\beta(\rho)}{\rho}\}}. \quad (2.10)$$

Denote by  $\mathbf{p}$  the duration of a busy period and by  $\mathbf{p}(\mathbf{v})$  that of an initial busy period, i.e.

$$\mathbf{p}(\mathbf{v}) := \inf_{t \geq 0} \{t : \mathbf{v}_t = 0 | \mathbf{v}_0 = \mathbf{v}\}, \quad \mathbf{v} \geq 0. \quad (2.11)$$

It is known, cf. [3], p. 261, formula (4.94), that: for  $\text{Re } \rho \geq 0, \mathbf{v} \geq 0$ ,

$$\mathbb{E}\{e^{-\rho \mathbf{P}(\mathbf{v})}\} := e^{-\frac{\rho \mathbf{v}}{1-a}} \mathbb{E}\{e^{-\rho \mathbf{P}_0}\}, \quad (2.12)$$

with  $\mathbf{p}_0$  a stochastic variable for which holds,

$$\mathbb{E}\{e^{-\rho \mathbf{P}_0}\} = 1-a + a \frac{1-\mathbb{E}\{e^{-\rho \mathbf{P}}\}}{\rho \mathbb{E}\{\mathbf{p}\}}, \quad \text{Re } \rho \geq 0, \quad (2.13)$$

note that

$$\mathbb{E}\{\mathbf{p}\} = \frac{1}{1-a}.$$

Since  $\mathbb{E}\{e^{-\rho \mathbf{P}}\}, \text{Re } \rho \geq 0$ , is for  $a < 1$  the unique zero of

$$\mu = \beta(\rho + a(1-\mu)), \quad |\mu| < 1, \quad (2.14)$$

a simple calculation shows that

$$\mathbb{E}\{e^{-\rho \mathbf{P}_0}\} = \mathbb{E}\{e^{-\rho \mathbf{P}(\mathbf{v})}\} = \mathbb{E}\{e^{-\frac{\rho}{1-a} \mathbf{v} \mathbb{E}\{e^{-\rho \mathbf{P}_0}\}}\}, \quad \text{Re } \rho \geq 0. \quad (2.15)$$

The probabilistic interpretation of the relation (2.15) is the following. For the stationary  $\mathbf{v}_t$ -process  $\mathbf{v}_t$  and  $\mathbf{v}$  have the same distribution. Hence  $\mathbf{p}_0 = 0$  if  $\mathbf{v}_t = 0$  and if  $\mathbf{v}_t > 0$  then  $\mathbf{p}_0$  has the distribution of the overshoot distribution of  $\mathbf{p}$ ; and consequently  $\mathbf{p}_0$  and  $\mathbf{p}(\mathbf{v})$  should have the same distribution.

From (2.15) it is seen that

$$z(\rho) := \mathbb{E}\{e^{-\rho \mathbf{P}_0}\}, \quad \text{Re } \rho \geq 0, \quad (2.16)$$

satisfies for  $a < 1$ ,

$$\begin{aligned} i. \quad & z = \mathbb{E}\{e^{-\frac{\rho}{1-a} \mathbf{V} z}\}, |z| \leq 1 \text{ for } \text{Re } \rho \geq 0, \\ ii \quad & z = 1 \text{ for } \rho = 0. \end{aligned} \quad (2.17)$$

It is simple to see that (2.17) has for  $\rho \geq 0$  a unique solution, and that it is positive. From

$$\frac{dz}{d\rho} [1-a + \mathbb{E}\{\mathbf{v} e^{-\frac{\rho}{1-a} \mathbf{V} z}\}] = -z \mathbb{E}\{e^{-\frac{\rho \mathbf{V}}{1-a} z}\}, \quad \rho \geq 0,$$

it is seen that

$$\frac{dz}{d\rho} \neq 0 \text{ for } \rho \geq 0.$$

This relation implies that this solution is regular for  $\rho > 0$ , continuous for  $\rho \geq 0$ . Consequently  $z(\rho), \rho \geq 0$ , cf. (2.16), is the unique solution of (2.17). Since the righthand side of (2.17) is regular for  $\operatorname{Re} \rho > 0$ , continuous for  $\operatorname{Re} \rho \geq 0$  whenever  $\operatorname{Re} z > 0$ , it follows by analytic continuation that  $z(\rho), \operatorname{Re} \rho \geq 0$ , is the unique solution of (2.17).

We quote some further relations from [3]. From [3], p. 260, formula (4.92), it is easily derived that: for  $\operatorname{Re} \rho \geq 0, v_0 \geq 0$ ,

$$\int_0^\infty e^{-\rho t} \Pr \{ \mathbf{v}_t = 0 | \mathbf{v}_0 = v_0 \} dt = \frac{1 - a e^{-\frac{\rho v_0}{1-a}} \mathbb{E} \{ e^{-\rho \mathbf{P}_0} \}}{\rho \mathbb{E} \{ e^{-\rho \mathbf{P}_0} \}}. \quad (2.18)$$

Further, from [3], p. 262, formula (4.99): for  $\operatorname{Re} \rho \geq 0, t \geq 0, v_0 \geq 0$ ,

$$\begin{aligned} \int_{0-}^\infty e^{-\rho \sigma} d \Pr \{ \mathbf{v}_t < \sigma | \mathbf{v}_0 = v_0 \} &= e^{-\rho v_0 + \rho t \{ 1 - a \frac{1 - \beta(\rho)}{\rho} \}} \\ -\rho U_1(t - v_0) \int_0^{t - v_0} e^{\rho \{ 1 - a \frac{1 - \beta(\rho)}{\rho} \} (t - v_0 - u)} \Pr \{ \mathbf{v}_{u + v_0} = 0 | \mathbf{v}_0 = v_0 \} du, \end{aligned} \quad (2.19)$$

with

$$U_1(x) = 0 \text{ for } x < 0, U_1(x) = 1 \text{ for } x \geq 0. \quad (2.20)$$

### 3. ON THE CONTRACTED VARIABLES

In this section we consider the contracted variables  $\Delta(a)\mathbf{v}, \Delta_1(a)\mathbf{p}(v)$  and  $\Delta_1(a)\mathbf{p}(\mathbf{v})$  for  $a \uparrow 1$ , here

$$\Delta_1(a) := \Delta(a)(1 - a). \quad (3.1)$$

From theorem 5.1 of [1] we have with  $\beta(\rho)$  as given in (2.6) for the M/G/1 model the following statement.

**THEOREM 3.1** *For the M/G/1 queue  $\Delta(a)\mathbf{v}$  converges in distribution for  $a \uparrow 1$ . The limiting distribution  $R_{\nu-1}(t)$  is given by*

$$\begin{aligned} R_{\nu-1}(t) &= \sum_{n=0}^\infty (-1)^n \frac{t^{n(\nu-1)}}{\Gamma(n(\nu-1) + 1)}, \quad t \geq 0, \\ &= \frac{1}{\pi} \sum_{n=1}^H (-1)^{n-1} \frac{\Gamma(n(\nu-1)) \sin n(\nu-1)\pi}{t^{n(\nu-1)}} + O(t^{-(H+1)(\nu-1)}) \text{ for } t \rightarrow \infty, \end{aligned} \quad (3.2)$$

for every finite  $H \in \{1, 2, \dots\}$ , and for  $\operatorname{Re} r \geq 0$ ,

$$\int_0^\infty e^{-rt} dR_{\nu-1}(t) = \lim_{a \uparrow 1} \mathbb{E} \{ e^{-r \Delta(a)\mathbf{v}} \} = \frac{1}{1 + r^{\nu-1}}. \quad (3.3)$$

**PROOF.** The theorem is a special case of that in [1], and so the proof here is slightly simpler. From (2.7)iii and (2.8) it follows that

$$g(\rho) = O(|\rho|) \text{ for } \operatorname{Re} \rho \geq 0, |\rho| \rightarrow 0. \quad (3.4)$$

From the definition of  $\Delta(a)$  it is readily seen that

$$\frac{\Delta(a)}{1-a} \rightarrow 0 \text{ for } a \uparrow 1. \quad (3.5)$$

Put in (2.10),

$$\rho = r\Delta(a), \text{ Re } r \geq 0,$$

so that we have from (2.6), (2.7), (2.10) and the definition of  $\Delta(a)$  that: for  $\text{Re } r \geq 0, 0 < 1 - a \ll 1$ ,

$$\begin{aligned} \mathbb{E}\{e^{-r\Delta(a)\mathbf{v}}\} &= \left[1 + \frac{a}{1-a} \left\{1 - \frac{1 - \beta(r\Delta(a))}{r\Delta(a)}\right\}\right]^{-1} \\ &= \left[1 + \frac{a}{1-a}g(r\Delta(a)) + \frac{ac}{1-a}\Delta(a)^{\nu-1}L(\Delta(a))r^{\nu-1}\frac{L(r\Delta(a))}{L(\Delta(a))}\right]^{-1} \\ &= \left[1 + \frac{a}{1-a}g(r\Delta(a)) + r^{\nu-1}\frac{L(r\Delta(a))}{L(\Delta(a))}\right]^{-1}. \end{aligned} \quad (3.6)$$

From (2.7)iv, (3.4) and (3.5) it is seen that: for  $a \uparrow 1$ ,

$$\frac{a}{1-a}g(r\Delta(a)) \rightarrow 0 \text{ and } \frac{L(r\Delta(a))}{L(\Delta(a))} \rightarrow 1, \text{ Re } r \geq 0, r \neq 0.$$

So the statement (3.3) follows. From this, from Feller's continuity theorem for LST's of probability distributions with support  $[0, \infty)$ , and from the continuity in  $r = 0$  of the last member of (3.3), it follows that  $\Delta(a)\mathbf{v}$  converges in distribution for  $a \uparrow 1$ . For the proof of the statements (3.2) concerning  $R_{\nu-1}(t)$  the reader is referred to [1].  $\square$

Next we consider the contracted variable

$$\Delta_1(a)\mathbf{p}_0,$$

with  $\Delta_1(a) = \Delta(a)(1-a)$ . Put

$$\rho = r\Delta_1(a), \text{ Re } r \geq 0, 0 < 1 - a \ll 1,$$

then it follows from (2.15) with

$$x(r; a) := \mathbb{E}\{e^{-r\Delta_1(a)\mathbf{p}_0}\}, \text{ Re } r \geq 0, 0 < 1 - a \ll 1, \quad (3.7)$$

that  $x(r, a)$  satisfies for  $0 < 1 - a \ll 1$ ,

$$\begin{aligned} i. \quad x &= \mathbb{E}\{e^{-r\Delta(a)\mathbf{v}x}\}, \text{ Re } r \geq 0, \\ ii. \quad x &= 1 \text{ for } r = 0. \end{aligned} \quad (3.8)$$

As in Section 2 it is shown that  $x(r; a), \text{ Re } r \geq 0$ , is the unique solution of (3.8). From Theorem 3.1 we know that  $\Delta(a)\mathbf{v}$  converges in distribution for  $a \uparrow 1$ . Let  $\mathbf{w}$  be a stochastic variable with distribution the limiting distribution  $R_{\nu-1}(t)$  of  $\Delta(a)\mathbf{v}$  for  $a \uparrow 1$ . As in Section 2 it is readily seen that the equation

$$\begin{aligned} i. \quad y &= \mathbb{E}\{e^{-r\mathbf{w}y}\}, \text{ Re } r \geq 0, \\ ii. \quad y &= 1 \quad \text{for } r = 0. \end{aligned} \quad (3.9)$$

has for  $r \geq 0$  a unique solution  $y(r)$ . This solution is real and regular for  $r > 0$ , continuous for  $r \geq 0$ ; note that (2.9)i implies that for  $r \geq 0$ ,

$$\frac{dy}{dr} \neq 0. \quad (3.10)$$



From (3.7) it is seen that the functions  $x(r; a)$ ,  $\text{Re } r \geq 0$  form a class of regular functions, which are uniformly bounded by one. Let  $a_n, n = 1, 2, \dots$ , with  $0 < 1 - a_n \ll 1$  be a sequence with  $a_n \uparrow 1$  for  $n \rightarrow \infty$ . From the regularity property just mentioned it follows that the sequence  $a_n, n = 1, 2, \dots$ , contains a subsequence  $a_{n_j}, j = 1, 2, \dots$ , such that the sequence  $x(r, a_{n_j}), \text{Re } r \geq 0$ , converges uniformly to a limit  $\tilde{y}(r)$  which is also a uniformly bounded function and which is regular for  $\text{Re } r > 0$ . cf. [4], p. 153. By dominated convergence it follows from (3.8) and (3.9) that  $\tilde{y}(r)$  satisfies (3.9). Because this conclusion holds for any sequence  $a_n, n = 1, 2, \dots$ , as defined above, it follows that the following limit

$$y(r) := \lim_{a \uparrow 1} x(r; a), \text{Re } r \geq 0, \quad (3.11)$$

exists, that it satisfies (3.9) and is regular for  $\text{Re } r > 0$ . Above it has been shown that the equation (3.9) has a unique solution for  $r \geq 0$ , and that it is regular for  $r > 0$ , and so by analytic continuation it is seen that  $y(r)$ , as defined in (3.11) is the unique solution of (3.9), and  $y(r)$  is regular for  $\text{Re } r > 0$ , and continuous for  $\text{Re } r \geq 0$ .

**THEOREM 3.2** *The stochastic variabel  $\Delta_1(a)\mathbf{p}_0$ , converges in distribution for  $a \uparrow 1$ , and its LST  $y(r), \text{Re } r \geq 0$ , is the unique solution of*

$$y = \frac{1}{1 + (ry)^{\nu-1}}, \quad \text{Re } r \geq 0, \quad (3.12)$$

$$y = 1 \quad \text{for } r = 0.$$

**PROOF.** Above it has been shown that  $y(r), \text{Re } r \geq 0$ , is the unique solution of (3.9), so by using Theorem 3.1 and by noticing that  $R_{\nu-1}(t)$  is the distribution of  $\mathbf{w}$  the statement (3.12) follows from (3.3) and (3.9). From (3.7) and (3.11) it follows that  $x(r, a)$ , the LST of  $\Delta_1(a)\mathbf{p}_0$ , converges for  $a \uparrow 1$  to  $y(r), \text{Re } r \geq 0$ . Because  $y(r)$  is continuous for  $r \geq 0$ , application of Feller's continuity theorem shows that  $\Delta_1(a)\mathbf{p}_0$  converges in distribution for  $a \uparrow 1$ .  $\square$

**COROLLARY 3.1** *The stochastic variable  $\Delta_1(a)(\mathbf{p}(\nu/\Delta(a))), \nu > 0$ , converges in distribution for  $a \uparrow 1$  and*

$$\lim_{a \uparrow 1} \mathbf{E}\{e^{-r\Delta_1(a)\mathbf{P}(\nu/\Delta(a))}\} = e^{-rvy(r)}, \quad \text{Re } r \geq 0, \nu > 0. \quad (3.13)$$

**PROOF.** From (2.12) we have with  $\rho = \Delta_1(a)$  for  $\text{Re } r \geq 0, 0 < 1 - a \ll 1$ , that

$$\mathbf{E}\{e^{-r\Delta_1(a)\mathbf{P}(\nu/\Delta(a))}\} = e^{-rv\mathbf{E}\{e^{-r\Delta_1(a)\mathbf{P}_0}\}}.$$

From this relation the proof follows similarly to that of the proof of theorem 3.2.  $\square$

**REMARK 3.1.** The solution  $y(r)$  of (3.12) is obviously a function of  $r^{\nu-1}$ . Put: for  $\text{Re } r \geq 0$ ,

$$z(r^{\nu-1}) := y(r), \quad (3.14)$$

so that  $z(s)$  is the unique solution of

$$z = \frac{1}{1 + sz^{\nu-1}}, \quad \text{Re } s \geq 0, \quad (3.15)$$

$$z = 1 \quad \text{for } s = 0.$$

Hence it follows for  $z(s)$  a solution of

$$z = 1 - sz^{\nu}, \quad (3.16)$$

that

$$\frac{dz}{ds}[1 - \nu s z^{\nu-1}] = -z^\nu,$$

So it is readily seen from

$$1 + \nu s_\nu z^{\nu-1}(s_\nu) = 0 \rightarrow z(s_\nu) = \frac{\nu}{\nu-1} \text{ and } s_\nu = -\frac{1}{\nu} \left(\frac{\nu-1}{\nu}\right)^{\nu-1}, \quad (3.17)$$

that  $s = s_\nu$  is the only branch point of the solution  $z(s)$  of (3.15). Since  $z(s)$  is regular for  $\operatorname{Re} s > 0$ , continuous for  $\operatorname{Re} s \geq 0$ , it follows that  $z(s)$  can be continued analytically from out  $\operatorname{Re} s \geq 0$  into the complex  $s$ -plane slitted along the halfline  $(-\infty, s_\nu)$ .

Denote by  $P_{\nu-1}(t)$  the limiting distribution of the distribution of  $\Delta_1(a)\mathbf{p}_0$  for  $a \uparrow 1$ . In appendix A the following asymptotic expression for  $P_{\nu-1}(t)$ ,  $t \rightarrow \infty$ , is derived.

For every finite  $H \in \{1, 2, \dots\}$  and  $t \rightarrow \infty$ ,

$$1 - P_{\nu-1}(t) = \sum_{n=1}^H (-1)^{n-1} t^{-n(\nu-1)} \frac{\Gamma(n\nu+1) \sin\{n(\nu-1)\pi\}}{n!n(\nu-1)(n(\nu-1)+1)} + O(t^{-(H+1)(\nu-1)}). \quad (3.18)$$

#### 4. ON THE CONTRACTED $\mathbf{v}_t$ -PROCESS

In this section we analyze the contracted  $\mathbf{v}_t$ -process. Before defining it we first consider the expression for

$$\omega(\rho, \theta) := \int_0^\infty e^{-\theta t} \mathbb{E}\{e^{-\rho \mathbf{V}_t} | \mathbf{v}_0 = 0\} dt, \quad (4.1)$$

with  $\operatorname{Re} \rho \geq 0$ ,  $\operatorname{Re} \theta > 0$ .

From (2.19) with  $\mathbf{v}_0 = 0$  we obtain: for  $\operatorname{Re} \rho = 0$ ,  $\operatorname{Re} \theta > 0$ ,

$$\omega(\rho, \theta) = \frac{1}{\theta - \rho \left\{1 - a \frac{1-\beta(\rho)}{\rho}\right\}} \left\{1 - \rho \int_0^\infty e^{-\theta t} \Pr\{\mathbf{v}_t = 0 | v_0 = 0\} dt\right\}, \quad (4.2)$$

with cf. (2.18): for  $\operatorname{Re} \theta \geq 0$ ,

$$\int_0^\infty e^{-\theta t} \Pr\{\mathbf{v}_t = 0 | \mathbf{v}_0 = 0\} dt = \frac{1-a}{\theta} \mathbb{E}^{-1}\{e^{-\theta \mathbf{P}_0}\}. \quad (4.3)$$

Put for  $0 < 1-a \ll 1$ ,

$$\begin{aligned} \rho &= r\Delta(a), \operatorname{Re} r \geq 0, t = \frac{\tau}{\Delta_1(a)}, \\ \theta &= s\Delta_1(a) = s\Delta(a)(1-a), \operatorname{Re} s \geq 0, \end{aligned} \quad (4.4)$$

and

$$\mathbf{w}_\tau(a) := \Delta(a)\mathbf{v}_{\tau/\Delta_1(a)}, \tau \geq 0. \quad (4.5)$$

The  $\{\mathbf{w}_\tau(a), \tau \geq 0\}$  process will be called the *contracted*  $\mathbf{v}_t$ -process. It follows that: for  $\operatorname{Re} r \geq 0$ ,  $\operatorname{Re} s > 0$ ,

$$\omega(r\Delta(a), s\Delta_1(a)) = \frac{1}{\Delta_1(a)} \int_0^\infty e^{-s\tau} \mathbb{E}\{e^{-r\mathbf{w}_\tau(a)}\} d\tau, \quad (4.6)$$

where we have deleted the conditioning event  $\mathbf{v}_0 = 0$  in the righthand side of (4.6). Put for  $\operatorname{Re} r \geq 0$ ,  $\operatorname{Re} s > 0$ ,  $0 < 1-a \ll 1$ ,

$$\Omega(r, s; a) := \int_0^{\infty} e^{-s\tau} \mathbb{E}\{e^{-r\mathbf{W}_\tau(a)}\} d\tau. \quad (4.7)$$

It follows from (4.2), (4.3) and (4.6) that: for  $\operatorname{Re} r \geq 0, \operatorname{Re} s > 0, 0 < 1 - a \ll 1$ ,

$$\Omega(r, s; a) = [s - r\{1 + \frac{a}{1-a}[1 - \frac{1 - \beta(r\Delta(a))}{r\Delta(a)}]\}]^{-1} [1 - \frac{r}{s} \mathbb{E}^{-1}\{e^{-s\Delta_1(a)} \mathbf{P}_0\}]. \quad (4.8)$$

In the proof of Theorem 2.1 it has been shown that

$$\frac{a}{1-a} \{1 - \frac{1 - \beta(r\Delta(a))}{r\Delta(a)}\} \rightarrow r^{\nu-1} \text{ for } a \uparrow 1 \text{ and } \operatorname{Re} r \geq 0. \quad (4.9)$$

From Theorem 3.1 we have: for  $\operatorname{Re} s \geq 0$ ,

$$\lim_{a \uparrow 1} \mathbb{E}\{e^{-s\Delta_1(a)} \mathbf{P}_0\} = y(s), \quad (4.10)$$

with  $y(r), \operatorname{Re} r \geq 0$ , the unique zero of (3.12).

From (4.8), (4.9), (4.10) and Theorem 3.2 we obtain: for  $\operatorname{Re} r \geq 0, \operatorname{Re} s > 0$ ,

$$\begin{aligned} \Omega(r, s) &:= \lim_{a \uparrow 1} \Omega(r, s; a) = \frac{1}{s - r(1 + r^{\nu-1})} \{1 - \frac{r}{s} y^{-1}(s)\} \\ &= \frac{1}{s} \frac{sy(s) - r}{s - r(1 + r^{\nu-1})} [1 + (sy(s))^{\nu-1}]. \end{aligned} \quad (4.11)$$

Note that the definition (4.7) of  $\Omega(r, s; a)$  implies that  $\Omega(r, s; a)$  is regular for  $\operatorname{Re} s > 0$ , since  $\mathbb{E}\{e^{-r\mathbf{W}_\tau(a)}\}$  is bounded for  $\operatorname{Re} r \geq 0$ .

Note also that by using (3.12)

$$s_0 = r_0(1 + r_0^{\nu-1}) \Leftrightarrow \frac{r_0}{s_0} = 1 + s_0^{\nu-1} (\frac{r_0}{s_0})^{\nu-1} \Leftrightarrow y(s_0) = \frac{r_0}{s_0}. \quad (4.12)$$

By using the inversion formula for the Laplace transformation we have from (4.7): for  $\operatorname{Re} r \geq 0, \varepsilon_1 > 0$ ,

$$\mathbb{E}\{e^{-r\mathbf{W}_\tau(a)}\} = \frac{1}{2\pi i} \int_{-i\infty + \varepsilon_1}^{i\infty + \varepsilon_1} e^{s\tau} \Omega(r, s; a) ds, \quad (4.13)$$

from which it follows by using (4.11) and dominated convergence that : for  $\operatorname{Re} r \geq 0, \varepsilon_1 > 0$ ,

$$\begin{aligned} \lim_{a \uparrow 1} \mathbb{E}\{e^{-r\mathbf{W}_\tau(a)}\} &= \frac{1}{2\pi i} \int_{-i\infty + \varepsilon_1}^{i\infty + \varepsilon_1} e^{s\tau} \Omega(r, s) ds \\ &= \frac{1}{2\pi i} \int_{-i\infty + \varepsilon_1}^{i\infty + \varepsilon_1} e^{st} \frac{1}{s - r(1 + r^{\nu-1})} \{1 - \frac{r}{s} y^{-1}(s)\} ds. \end{aligned} \quad (4.14)$$

The righthand side of (4.14) is obviously a continuous function of  $r$  for  $r \geq 0$ , and it is seen that it converges for  $r \downarrow 0$  to

$$\frac{1}{2\pi i} \int_{-i\infty + \varepsilon_1}^{i\infty + \varepsilon_1} e^{s\tau} \frac{ds}{s} = 1 \text{ for } \tau > 0.$$

Hence it follows from Feller's continuity theorem for the LST of a distribution with support contained in  $[0, \infty)$  that

$\mathbf{w}_\tau(a)$  converges for  $a \uparrow 1$  in distribution for every  $\tau > 0$ . (4.15)

Denote by  $\mathbf{w}_\tau$  a stochastic variable with distribution the limiting distribution of  $\mathbf{w}_\tau(a)$  for  $a \uparrow 1$ .

From (4.7), (4.11) and (4.15) we then obtain by using dominated convergence that: for  $\operatorname{Re} r \geq 0$ ,  $\operatorname{Re} s > 0$ ,

$$\begin{aligned} \int_0^\infty e^{-s\tau} \mathbb{E}\{e^{-r\mathbf{W}_\tau}\} d\tau &= \Omega(r, s) = \\ &= \frac{1}{s} \frac{sy(s) - r}{s - r(1 + r^{\nu-1})} [1 + (sy(s))^{\nu-1}]. \end{aligned} \quad (4.16)$$

Because  $y(0) = 1$ , see Theorem 3.2, we obtain

$$\lim_{s \downarrow 0} \int_0^\infty se^{-s\tau} \mathbb{E}\{e^{-r\mathbf{W}_\tau}\} d\tau = \frac{1}{1 + r^{\nu-1}} \text{ for } \operatorname{Re} r \geq 0. \quad (4.17)$$

From (4.5) and (4.15) we have: for  $\operatorname{Re} r \geq 0$ ,  $\tau > 0$ ,

$$\mathbb{E}\{e^{-r\mathbf{W}_\tau}\} = \lim_{a \uparrow 1} \mathbb{E}\{e^{-r\mathbf{W}_\tau(a)}\} = \lim_{a \uparrow 1} \mathbb{E}\{e^{-r\Delta(a)\mathbf{V}_\tau/\Delta_1(a)}\},$$

from which it follows, since uniform convergence and Theorem 3.1 imply, that for  $\operatorname{Re} r \geq 0$ ,

$$\begin{aligned} \lim_{\tau \uparrow \infty} \lim_{a \uparrow 1} \mathbb{E}\{e^{-r\Delta(a)\mathbf{V}_\tau/\Delta_1(a)}\} &= \lim_{a \uparrow 1} \mathbb{E}\{e^{-r\Delta(a)\mathbf{V}}\} = \frac{1}{1 + r^{\nu-1}}, \\ \lim_{\tau \rightarrow \infty} \mathbb{E}\{e^{-r\mathbf{W}_\tau}\} &= \frac{1}{1 + r^{\nu-1}}. \end{aligned} \quad (4.18)$$

This result also follows from (4.17) by using a wellknown Abel theorem for the Laplace transform when it is known that  $\mathbf{w}_\tau$  converges in distribution for  $\tau \rightarrow \infty$ . From (4.3), (4.4) and (4.5) we obtain: for  $\operatorname{Re} s > 0$ ,  $0 < 1 - a \ll 1$ ,

$$\int_0^\infty e^{-st} \Pr\{\mathbf{w}_t(a) = 0 | \mathbf{v}_0 = 0\} dt = \frac{1-a}{s} \mathbb{E}^{-1}\{e^{-s\Delta_1(a)\mathbf{P}_0}\}. \quad (4.19)$$

Hence, by using the inversion formula for the Laplace Transform, we have: for  $t > 0$ ,  $0 < 1 - a \ll 1$ , and  $\varepsilon > 0$ ,

$$\Pr\{\mathbf{w}_\tau(a) = 0 | \mathbf{v}_0 = 0\} = \frac{1-a}{2\pi i} \int_{-i\infty+\varepsilon}^{i\infty+\varepsilon} \frac{e^{s\tau}}{s} \mathbb{E}^{-1}\{e^{-s\Delta_1(a)\mathbf{P}_0}\} ds, \quad (4.20)$$

From Theorem (3.2) we have: for  $\operatorname{Re} s \geq 0$ ,

$$\lim_{a \uparrow 1} \mathbb{E}^{-1}\{e^{-s\Delta_1(a)\mathbf{P}_0}\} = y^{-1}(s),$$

note that (3.12) implies that  $y(s) \neq 0$  for  $\operatorname{Re} s \geq 0$ . Hence from (4.13) by using dominated convergence it is seen that for  $\tau \geq 0$ , the following limit exists and

$$\lim_{a \uparrow 1} \frac{1}{a-1} \Pr\{\mathbf{w}_\tau(a) = 0 | \mathbf{v}_0 = 0\} = \frac{1}{2\pi i} \int_{-i\infty+\varepsilon}^{i\infty+\varepsilon} \frac{e^{s\tau}}{s} \frac{ds}{y(s)}. \quad (4.21)$$

From the results derived above it is seen that the following theorem results.

**THEOREM 4.1.** *For the M/G/1 queueing model with traffic load  $a < 1$ , with message length distribution characterized by (3.6) and  $\Delta(a)$  that zero of  $ax^{\nu-1}L(x)/(1-a) = 1, x > 0$ , which tends to zero for  $a \uparrow 1$ , holds for the workload  $\mathbf{v}_t, t \geq 0$  that:*

- i. the contracted workload  $\mathbf{w}_\tau(a) := \Delta_1(a)\mathbf{v}_{\tau/\Delta_1(a)}$  with  $\Delta_1(a) = \Delta(a)(1-a)$  for  $0 < (1-a)/ac < 1$  converges in distribution for  $a \uparrow 1$  and every  $\tau > 0$ ;*
- ii. with  $\mathbf{w}_\tau$  a stochastic variable with distribution the limiting distribution of  $\mathbf{w}_\tau(a)$  for  $a \uparrow 1$ , holds:*  
for  $\text{Re } r \geq 0, \tau \geq 0, \varepsilon > 0$ ,

$$\mathbb{E}\{e^{-r\mathbf{w}_\tau}\} = \frac{1}{2\pi i} \int_{-i\infty+\varepsilon}^{i\infty+\varepsilon} \frac{e^{s\tau}}{s-r(1+r^{\nu-1})} \left\{1 - \frac{r}{s}y^{-1}(s)\right\} ds, \quad (4.22)$$

here  $y(s)$  is the LST of  $\lim_{a \uparrow 1} \Delta_1(a)\mathbf{p}_0$ , cf. Theorem 3.2;

- iii. for  $\text{Re } r \geq 0$ ,*

$$\lim_{\tau \rightarrow \infty} \mathbb{E}\{e^{-r\mathbf{w}_\tau}\} = \frac{1}{1+r^{\nu-1}},$$

and  $\mathbf{w}_\tau$  for  $\tau \rightarrow \infty$  and  $\Delta(a)\mathbf{v}$  for  $a \uparrow 1$  have the same limiting distribution  $R_{\nu-1}(t)$ , see Theorem 2.1,

- iv. for  $\text{Re } r \geq 0, \text{Re } s > 0$ ,*

$$\int_0^\infty e^{-s\tau} \mathbb{E}\{e^{-r\mathbf{w}_\tau}\} d\tau = \frac{1}{s-r(1+r^{\nu-1})} \left\{1 - \frac{r}{s}y^{-1}(s)\right\}.$$

**REMARK 4.1.** The statements have actually been proved for the case that  $\mathbf{v}_0 = 0$ , i.e.  $\mathbf{w}_0 = 0$ , so that the expectations in (3.24) should be conditional expectations, the conditional event being  $\mathbf{w}_0 = 0$ . However, the analysis above needs hardly any change whenever  $\mathbf{v}_0$  or  $\mathbf{w}_0$  are positive, and it is then seen that the resulting limiting distributions are independent of  $\mathbf{w}_0$ .  $\square$

**REMARK 4.2.** In the relations (4.24) the average message length distribution  $\beta$  has been taken as the time unit, see below (2.1). When we do not make this convention then  $\tau$  should be replaced by  $\tau/\beta$  and  $\mathbf{w}_\tau$  by  $\mathbf{w}_{\tau/\beta}/\beta$ .  $\square$

## 5. THE CONTRACTED INPUT PROCESS

In this section we consider the input process  $\{\mathbf{k}_t, t \geq 0\}$  of the M/G/1-queue. Here  $\mathbf{k}_t$  is the total amount of work generated by the arrivals in  $[0, t)$ ,  $\mathbf{k}_0 = 0$ . The virtual backlog  $\mathbf{h}_t$  and the noise-traffic  $\mathbf{n}_t$  at time  $t$  are defined by

$$\begin{aligned} \mathbf{h}_t &= \mathbf{k}_t - t, \\ \mathbf{n}_t &= \mathbf{k}_t - at. \end{aligned} \quad (5.1)$$

The virtual waiting time or workload  $\mathbf{v}_t$  at time  $t$  is given by Reich's formula, see [3], p. 170: with  $\mathbf{v}_0 = \mathbf{v}_0 \geq 0$ ,

$$\mathbf{v}_t = \max[\mathbf{v}_0 + \mathbf{h}_t, \sup_{0 < u < t} (\mathbf{h}_t - \mathbf{h}_u)]. \quad (5.2)$$

It is wellknown that: for  $\text{Re } \rho \geq 0, t \geq 0$ ,

$$\mathbf{E}\{e^{-\rho \mathbf{k}_t}\} = \sum_{n=0}^{\infty} \frac{(at)^n}{n!} e^{-at} \beta^n(\rho) = e^{-a\rho t \frac{1-\beta(\rho)}{\rho}}. \quad (5.3)$$

From the last three relations it follows that: for  $\operatorname{Re} \rho = 0, t \geq 0$ ,

$$\begin{aligned} \mathbf{E}\{e^{-\rho \mathbf{n}_t}\} &= e^{a\rho t \{1 - \frac{1-\beta(\rho)}{\rho}\}}, \\ \mathbf{E}\{e^{-\rho \mathbf{h}_t}\} &= e^{\rho t \{1 - a \frac{1-\beta(\rho)}{\rho}\}}. \end{aligned} \quad (5.4)$$

Note that

$$\mathbf{E}\{\mathbf{n}_t\} = \mathbf{E}\{\mathbf{k}_t - at\} = 0, \quad t \geq 0. \quad (5.5)$$

We introduce the following scaling, cf. (4.4).

$$\begin{aligned} \rho &= r\Delta(a), \quad \operatorname{Re} r \geq 0, \\ t &= \tau/\Delta_1(a). \end{aligned} \quad (5.6)$$

With this scaling we define the contracted versions of  $\mathbf{n}_t, \mathbf{h}_t$  and  $\mathbf{v}_t$ ,

$$\begin{aligned} \mathbf{N}(\tau; a) &:= \Delta(a) \mathbf{n}_{\tau/\Delta_1(a)}, \\ \mathbf{H}(\tau; a) &:= \Delta(a) \mathbf{h}_{\tau/\Delta_1(a)}, \\ \mathbf{w}_{\tau}(a) &:= \Delta(a) \mathbf{v}_{\tau\Delta_1(a)}. \end{aligned} \quad (5.7)$$

It follows from (5.4) and (5.7) that: for  $0 < 1 - a \ll 1, \operatorname{Re} r = 0, \tau \geq 0$ ,

$$\begin{aligned} \mathbf{E}\{e^{-r\mathbf{N}(\tau; a)}\} &= e^{\frac{ar\tau}{1-a} \{1 - \frac{1-\beta(r\Delta(a))}{r\Delta(a)}\}}, \\ \mathbf{E}\{e^{-r\mathbf{H}(\tau; a)}\} &= e^{\frac{r\tau}{1-a} \{1 - a \frac{1-\beta(r\Delta(a))}{r\Delta(a)}\}}, \\ \mathbf{H}(\tau; a) &= \mathbf{N}(\tau; a) - \tau; \end{aligned} \quad (5.8)$$

and from (5.3),

$$\mathbf{w}_{\tau}(a) = \max[\Delta(a) \mathbf{v}_0 + \mathbf{H}(\tau; a), \sup_{0 < u < \tau} (\mathbf{H}(\tau; a) - \mathbf{H}(u; a))]. \quad (5.9)$$

**THEOREM 5.1.**  $\mathbf{N}(\tau; a)$  and  $\mathbf{H}(\tau; a)$  both converge in distribution for  $a \uparrow 1$  for every  $\tau > 0$  and: for  $\operatorname{Re} r = 0$ ,

$$\begin{aligned} i. \quad \lim_{a \uparrow 1} \mathbf{E}\{e^{-r\mathbf{N}(\tau; a)}\} &= e^{\tau r^{\nu}}, \\ ii. \quad \lim_{a \uparrow 1} \mathbf{E}\{e^{-r\mathbf{H}(\tau; a)}\} &= e^{r\tau(1+r^{\nu-1})}. \end{aligned} \quad (5.10)$$

**PROOF.** The relations (5.10) follow from (5.8) by using the same arguments as in the proof of Theorem 3.1. Because the righthand sides in (5.10) are both continuous at  $r = 0$ , it follows from the continuity theorem for characteristic functions that  $\mathbf{N}(\tau; a)$ , and similarly  $\mathbf{H}(\tau; a)$ , converges in distribution for  $a \uparrow 1$ .  $\square$

Let  $\mathbf{N}(\tau)$  and  $\mathbf{H}(\tau) = \mathbf{N}(\tau) - \tau$ , be stochastic variables for which holds: for  $\operatorname{Re} r = 0, \tau \geq 0, 1 < \nu \leq 2$ ,

$$\begin{aligned} \mathbf{E}\{e^{-r\mathbf{N}_{tau}}\} &= e^{\tau r^{\nu}}, \\ \mathbf{E}\{e^{-r\mathbf{H}(\tau)}\} &= e^{r\tau(1+r^{\nu-1})} = e^{r\tau} \mathbf{E}\{e^{-r\mathbf{N}(\tau)}\}. \end{aligned} \quad (5.11)$$

The distributions of  $\mathbf{N}(\tau)$  and  $\mathbf{H}(\tau)$  may be characterized by using the notation in [5], p. 11. Here  $\mathcal{S}_\nu(\sigma, 1, \mu)$  stands for the distribution of a stochastic variable  $\mathbf{x}$  with characteristic function:

$$\mathbb{E}\{e^{i\theta\mathbf{x}}\} = e^{-\sigma^\nu |\theta|^\nu \{1 - i \frac{\theta}{|\theta|} \tan \frac{1}{2}\nu\pi\} + i\mu\theta}, \quad (5.12)$$

with

$$\sigma \geq 0, 1 < \nu \leq 2, \mu \text{ and } \theta \text{ both real.} \quad (5.13)$$

The distribution  $\mathcal{S}_\nu(\sigma, 1, \mu)$  belongs to the class of stable distributions. From (5.4) and (5.12) it follows, cf. [5], p. 15 and 51, that: for  $\tau \geq 0, 1 < \nu \leq 2$ ,

$$\begin{aligned} \mathcal{S}_\nu([-\tau \cos \frac{1}{2}\nu\pi]^{1/\nu}, 1, 0) & \text{ is the distribution of } \mathbf{N}(\tau), \\ \mathcal{S}_\nu([-\tau \cos \frac{1}{2}\nu\pi]^{1/\nu}, 1, -\tau) & \text{ is the distribution of } \mathbf{H}(\tau). \end{aligned} \quad (5.14)$$

Note that (5.11) implies that  $\mathbf{N}(\tau)$  and  $\sigma^{1/\nu}\mathbf{N}(\tau/\sigma), \sigma > 0$ , have the same distribution, i.e.  $\mathbf{N}(\tau)$  is self-similar with index  $1/\nu$ .

From the definition of  $\mathbf{N}(\tau)$  it is readily seen that the process  $\{\mathbf{N}(\tau), \tau \geq 0\}$  has stationary independent increments because

$$\mathbf{N}(\tau_2) - \mathbf{N}(\tau_1) \text{ and } \mathbf{N}(\tau_4) - \mathbf{N}(\tau_3)$$

are independent for  $0 \leq \tau_1 < \tau_2 \leq \tau_3 < \tau_4$ , and

$$\mathbf{N}(\tau_2) - \mathbf{N}(\tau_1) \stackrel{d}{=} \mathbf{N}(\tau_2 - \tau_1), \tau_2 > \tau_1 \geq 0, \quad (5.15)$$

here  $\stackrel{d}{=}$  stands for ‘‘have the same distribution’’.

It is readily seen, cf. [5], p.113 and 349, that the  $\mathbf{N}(\tau)$  process is a  $\nu$ -stable Lévy motion for  $1 < \nu < 2$ , for  $\nu = 2$  it is Brownian motion.

The same arguments show that the process  $\{\mathbf{H}(\tau), \tau \geq 0\}$  has stationary, independent increments.

Next, we introduce the contracted M/G/1 model. Its workload process is the process  $\{\mathbf{w}_\tau(a), \tau \geq 0\}$  with  $\mathbf{w}_\tau(a)$  as given by (5.9). Its virtual backlog is given by  $\mathbf{H}(\tau; a)$  and its noise traffic by  $\mathbf{N}(\tau; a)$ . From (5.1) and (5.7) we have for  $0 < a < 1, \tau \geq 0$ ,

$$\mathbb{E}\{\mathbf{N}(\tau; a)\} = 0, \quad \mathbb{E}\{\mathbf{H}(\tau; a)\} = -\tau, \quad (5.16)$$

$$\mathbf{H}(\tau; a) = \mathbf{N}(\tau; a) - \tau.$$

We next consider the  $\tilde{M}/\tilde{G}/1$  model with noise traffic  $\mathbf{N}(\tau)$ , virtual backlog  $\mathbf{H}(\tau) = \mathbf{N}(\tau) - \tau$  and workload  $\tilde{\mathbf{w}}_\tau$  at time  $\tau$  defined by

$$\tilde{\mathbf{w}}_\tau = \max[\mathbf{H}(\tau), \sup_{0 < u < \tau} (\mathbf{H}(\tau) - \mathbf{H}(u))]. \quad (5.17)$$

The question arises whether this  $\tilde{M}/\tilde{G}/1$  model can be considered as the ‘‘limit’’ for  $a \uparrow 1$  of the above described contracted M/G/1 model, i.e. whether the  $\{\mathbf{w}_\tau(a), \tau \geq 0\}$  process converges weakly for  $a \uparrow 1$  to the  $\{\tilde{\mathbf{w}}_\tau, \tau \geq 0\}$  processes. This is indeed the case and we now sketch its proof. It has been shown above that the  $\{\mathbf{N}(\tau; a), \tau > 0\}$  process has homogeneous independent increments, i.e. the distribution of  $\mathbf{N}(\tau_1; a) - \mathbf{N}(\tau_2; a)$  is a function of  $\tau_2 - \tau_1$ . From the results above it is readily seen that the finite dimensional distributions of the increments of the  $\{\mathbf{N}(\tau; a), \tau > 0\}$  process converge weakly to those of the  $\{\mathbf{N}(\tau), \tau > 0\}$  process, which is also a process with homogeneous independent increments. The analogous statements hold for the  $\mathbf{H}(\tau; a)$  and the  $\mathbf{H}(\tau)$ -process. The functional  $\sup_{0 < u < \tau} \{\mathbf{H}(\tau; a) - \mathbf{H}(u; a)\}$  is a continuous functional of the  $\mathbf{H}(\tau; a)$  process and it satisfies the conditions of Corollary 3.2 of [11]. It is readily verified that this corollary, when applied to the  $\mathbf{H}(\tau; a)$

and  $\mathbf{H}(\tau)$  process, shows that the just mentioned functional converges for  $a \uparrow 1$  in the Skorokhod topology to the analogous functional of the  $\mathbf{H}(\tau)$ -process. It then follows from (5.9) and (5.17) that the  $\{\mathbf{w}(\tau; a), \tau > 0\}$  process converges in the Skorokhod topology to the  $\{\tilde{\mathbf{w}}(\tau), \tau > 0\}$  process.

In Section 4 it has been shown that  $\mathbf{w}_\tau(a)$  converges in distribution for  $a \uparrow 1$ , so that from the definition of  $\mathbf{w}_\tau, \tau \geq 0$ , cf. (4.22)i, it is seen that

$$\tilde{\mathbf{w}}_\tau \stackrel{d}{=} \mathbf{w}_\tau. \quad (5.18)$$

Consequently, the statements (4.22)ii, iii and iv also hold for  $\tilde{\mathbf{w}}_\tau$ .

By using the notation

$$\mathbf{y}_u \xrightarrow[u \rightarrow u_0]{d} \mathbf{y}$$

to express that  $\mathbf{y}_u$  converges in distribution for  $u \rightarrow u_0$  with limiting distribution that of  $\mathbf{y}$ , the limit results obtained above may be written as follows.

From Theorem 3.1, from (4.5) and (4.24)iii and  $1 < \nu \leq 2$ ,

$$\mathbf{w}_\tau(a) = \Delta(a)\mathbf{v}_{\tau/\Delta_1(a)} \xrightarrow[a \uparrow 1]{d} \mathbf{w}_\tau \xrightarrow[\tau \rightarrow \infty]{d} \mathbf{w}_\infty, \quad (5.19)$$

$$\mathbf{w}_\tau(a) \xrightarrow[\tau \rightarrow \infty]{d} \Delta(a)\mathbf{v} \xrightarrow[a \uparrow 1]{d} \mathbf{w}_\infty,$$

and for  $1 < \nu < 2$ ,

$$\mathbf{w}_\tau(a) \xrightarrow[a \uparrow 1]{d} \tilde{\mathbf{w}}_\tau \xrightarrow[\tau \rightarrow \infty]{d} \mathbf{w}_\infty, \quad (5.20)$$

here  $\mathbf{w}_\infty$  has the distribution  $R_{\nu-1}(t)$ , cf. (3.2).

#### APPENDIX A

In this Appendix we derive an asymptotic expression for the limiting distribution  $P_{\nu-1}(t)$  of the distribution of  $\Delta_1(a)\mathbf{p}_0$  for  $a \uparrow 1$ , cf. Theorem 3.2.

With  $z(s)$  as defined in Remark 3.1 put

$$x(s) := z(s) - 1, \quad (a.1)$$

so that

$$x(s) = (-s)(1 + x(s))^\nu, \quad (a.2)$$

$$x(0) = 0.$$

Let  $C$  be the circle with center at  $x = 0$  and radius  $x_\nu$ , with

$$x_\nu := x(s_\nu) = z(s_\nu) - 1 = \frac{1}{\nu - 1}, \quad (a.3)$$

and  $s_\nu$  as defined in (3.17). It is readily seen that: for  $|s| < |s_\nu|$ ,



$$|s||1+x|^\nu < |x| \text{ for } |x| < |x_\nu|. \quad (\text{a.4})$$

Because  $x$  and  $1+x$  are both regular functions of  $x$  for  $|x| < |s_\nu|$ , it follows from Rouché's theorem that the equation  $x = (-s)(1+x)^\nu$  has a unique root inside the circle with center at  $x = 0$  and radius  $|s_\nu| - \varepsilon$  for  $0 < \varepsilon \ll 1$ . Hence, we can apply Langrange's theorem [10], p.133, for the derivation of a power series for  $x(s)$ . It results that: for  $|s| < |s_\nu|$ ,

$$x(s) = \sum_{n=0}^{\infty} (-1)^n \frac{d^{n-1}}{dt^{n-1}} (1+t)^{n\nu} |_{t=0} = \sum_{n=1}^{\infty} (-1)^n s^n \frac{\Gamma(n\nu+1)}{n! \Gamma(n(\nu-1)+2)}.$$

Hence from (3.14), (a.1) and the definition of  $y(r)$ , cf. Theorem 3.2, we have

$$y(r) = 1 + \sum_{n=1}^{\infty} (-1)^n r^{n(\nu-1)} \frac{\Gamma(n\nu+1)}{n! \Gamma(n(\nu-1)+2)}, \quad (\text{a.5})$$

for

$$0 \leq \operatorname{Re} r^{\nu-1} < -s_\nu,$$

or

$$\frac{1-y(r)}{r} = \sum_{n=1}^{\infty} (-1)^{n-1} r^{n(\nu-1)-1} \frac{\Gamma(n\nu+1)}{n! \Gamma(n(\nu-1)+2)}. \quad (\text{a.6})$$

Since  $y(r)$  is the LST of  $P_{\nu-1}(t)$ , we have: for  $\operatorname{Re} r \geq 0$ ,

$$\frac{1-y(r)}{r} = \int_0^{\infty} e^{-rt} \{1 - P_{\nu-1}(t)\} dt. \quad (\text{a.7})$$

From (a.6) and (a.7) we derive the asymptotic relation for  $1 - P_{\nu-1}(t) \rightarrow \infty$ , by using Theorem 2 of [2], vol.II. p.159. The function  $y(r)$  is regular for  $\operatorname{Re} r > 0$ , continuous for  $\operatorname{Re} r \geq 0$  and  $r = 0$  is a branch point of  $y(r)$ . From (3.12) it is seen that  $y(r)$  can be continued analytically into  $\{r : |r| > 0, -\frac{1}{2}\pi - \omega < \arg r < \frac{1}{2}\pi + \omega\}$  for a  $\omega \in (0, \frac{1}{2}\pi)$  and that  $y(r)$  is absolutely integrable on every finite interval with  $|\arg r| < u$ ,  $|u| < \frac{1}{2}\pi + \omega$ . From the just mentioned Theorem of [2] and from (a.6) and (a.7) it follows that: for  $t \rightarrow \infty$  and every finite  $H \in \{1, 2, \dots\}$ ,

$$1 - P_{\nu-1}(t) = \sum_{n=1}^H (-1)^{n-1} t^{-n(\nu-1)} \frac{\Gamma(n\nu+1)}{n! \Gamma(n(\nu-1)+2)} \frac{1}{\Gamma(1-n(\nu-1))} + O(t^{-(H+1)(\nu-1)}). \quad (\text{a.8})$$

Using the relation

$$\Gamma^{-1}(1-z) = \Gamma(z) \frac{\sin \pi z}{z}$$

it is seen that (a.8) may be rewritten as:

$$1 - P_{\nu-1}(t) = \frac{1}{\pi} \sum_{n=1}^H (-1)^{n-1} t^{-n(\nu-1)} \frac{\Gamma(n\nu+1) \sin \{n(\nu-1)\pi\}}{n! n(\nu-1)(n(\nu-1)+1)} + O(t^{-(H+1)(\nu-1)}) \text{ for } t \rightarrow \infty, \quad (\text{a.9})$$

and so (a.9) is the asymptotic expression for the tail of the distribution  $P_{\nu-1}(t)$ .

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