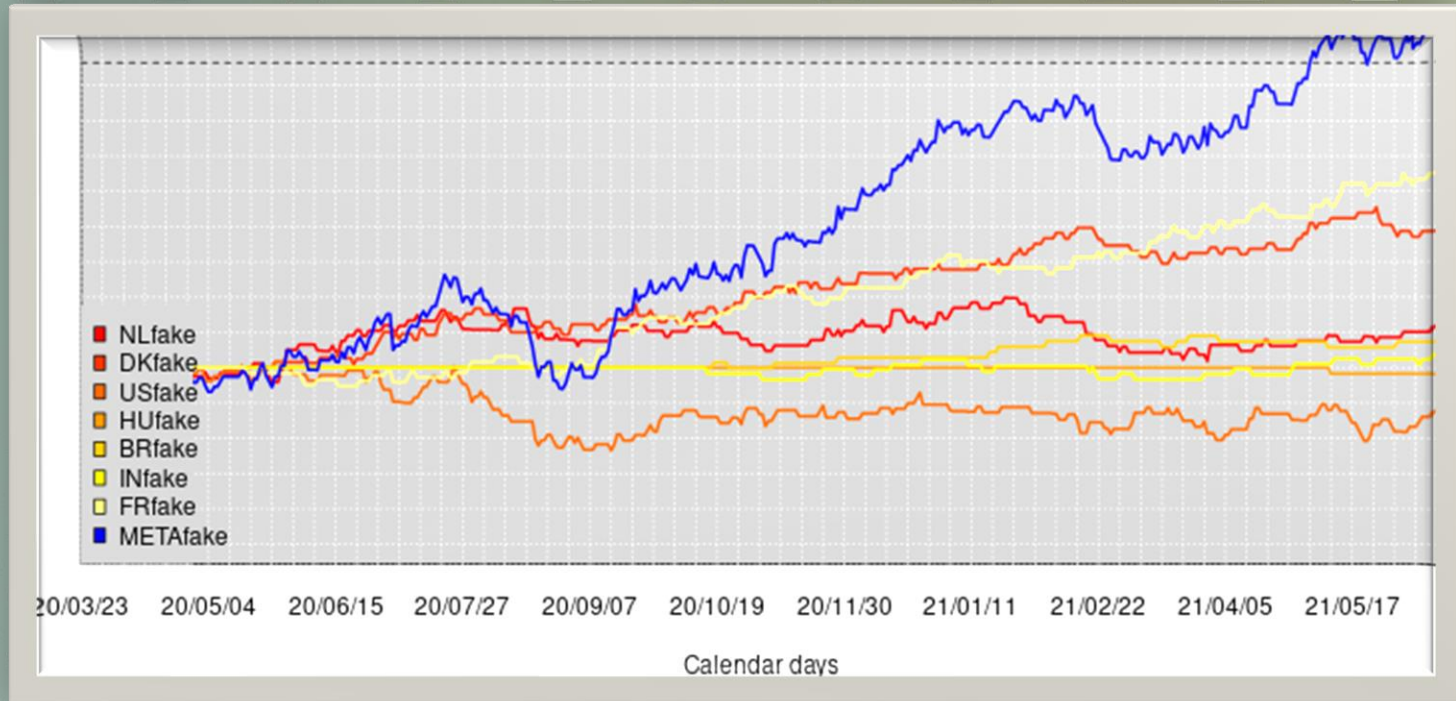


ALL-IN meta-analysis



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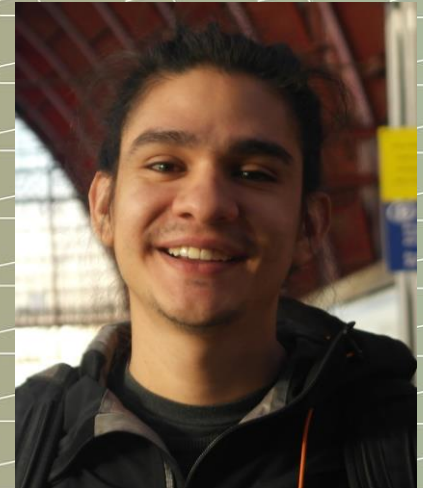
Peter Grünwald



Judith ter Schure



Alexander Ly



Muriel Pérez

ALL-IN meta-analysis properties

Type-I error whenever you analyze

- Continuous monitoring
- Unlimited horizon (no prespecified maximum sample size)
- Possibly 100% power
- Live meta-analysis (update after each new observation)
- Effortless cumulative meta-analysis (unlimited number of trials)

ALL-IN meta-analysis

Anytime

Live and

Leading

***IN*terim meta-analysis**

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I'll give an example of time-to-event analysis (logrank test)
ALL-IN-META-BCG-CORONA

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ALL-IN-META-BCG-CORONA

Whenever you see e-value, you can read: likelihood ratio

ALL-IN meta-analysis

- Meta evidence by *multiplying* study e-values
- **Type-I error control** under continuous monitoring and **any decision** to start, stop or expand studies
- Evidence against **the global null** (= no risk difference *in any study*) contribution by any study with hazard ratio < 0.8

e-values and Ville's inequality

$$E\langle i \rangle = \frac{P(\text{observed event } i \mid \text{hazard ratio} = 0,8)}{P(\text{observed event } i \mid \text{hazard ratio} = 1)} \frac{H_1}{H_0}$$

- If $E\langle i \rangle$ is an *e-value* for a single event i

$$E\langle n \rangle = \prod_{i=1}^n E\langle i \rangle$$

is also an *e-value* for n events

- Ville's inequality (1939)

$$P_{\text{events} \sim H_0} \left(E\langle n \rangle \geq \frac{1}{\alpha} \text{ for some } n \right) \leq \alpha$$

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$$E\langle i \rangle = \frac{P(\text{observed event } i \mid \text{hazard ratio} = 0,8)}{P(\text{observed event } i \mid \text{hazard ratio} = 1)} \quad \begin{matrix} H_1 \\ H_0 \end{matrix}$$

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e-values and Ville's inequality

$$E^{\langle n \rangle} = \prod_{i=1}^n \frac{\mathbf{P}(\text{observed event } i \mid \text{hazard ratio} = 0,8)}{\mathbf{P}(\text{observed event } i \mid \text{hazard ratio} = 1)} \frac{H_1}{H_0}$$

- If $E_1^{\langle n_1(t) \rangle}$ is an *e-value* for n_1 events at time t in study 1, testing H_0
 \vdots
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e-values and Ville's inequality

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 $E_K^{\langle n_K(t) \rangle}$ is an *e-value* for n_K events at time t in study K , testing H_0

$$E^{\langle t \rangle} = \prod_{k=1}^K E_k^{\langle n_k(t) \rangle}$$

is also an *e-value* for $n_1(t) + \dots + n_K(t)$ events at time t , testing H_0

e-values and Ville's inequality

$$E^{(n)} = \prod_{i=1}^n \frac{\mathbf{P}(\text{observed event } i \mid \text{hazard ratio} = 0,8)}{\mathbf{P}(\text{observed event } i \mid \text{hazard ratio} = 1)} \frac{H_1}{H_0}$$

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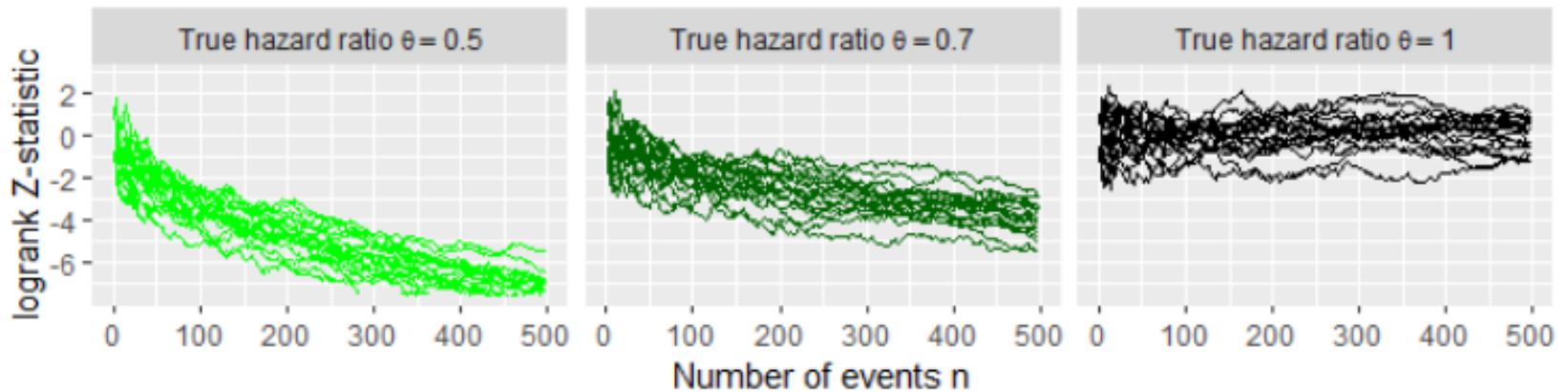
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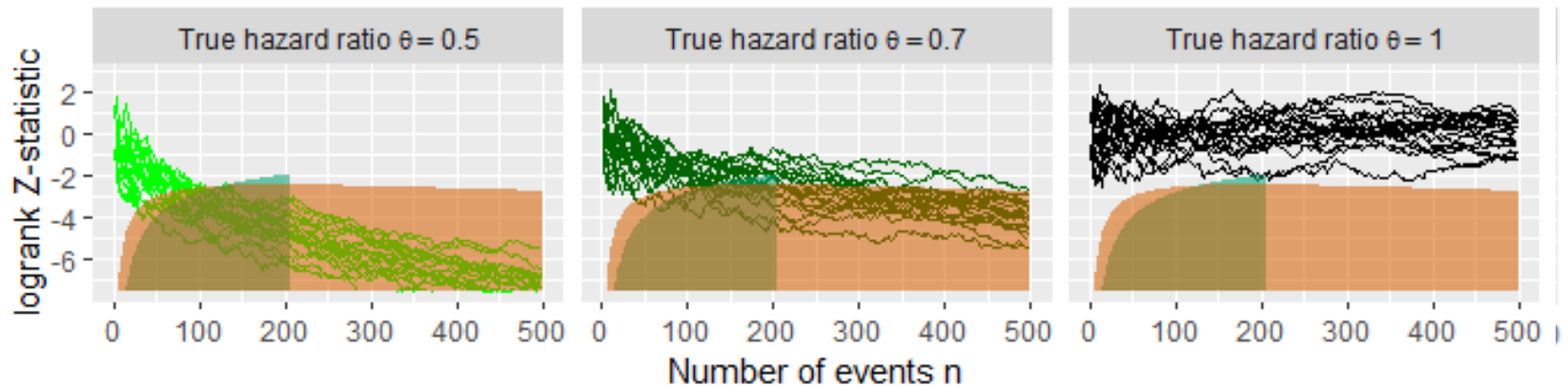
Possibly 100% power

Null hypothesis rejections by simulated data for a left-sided test



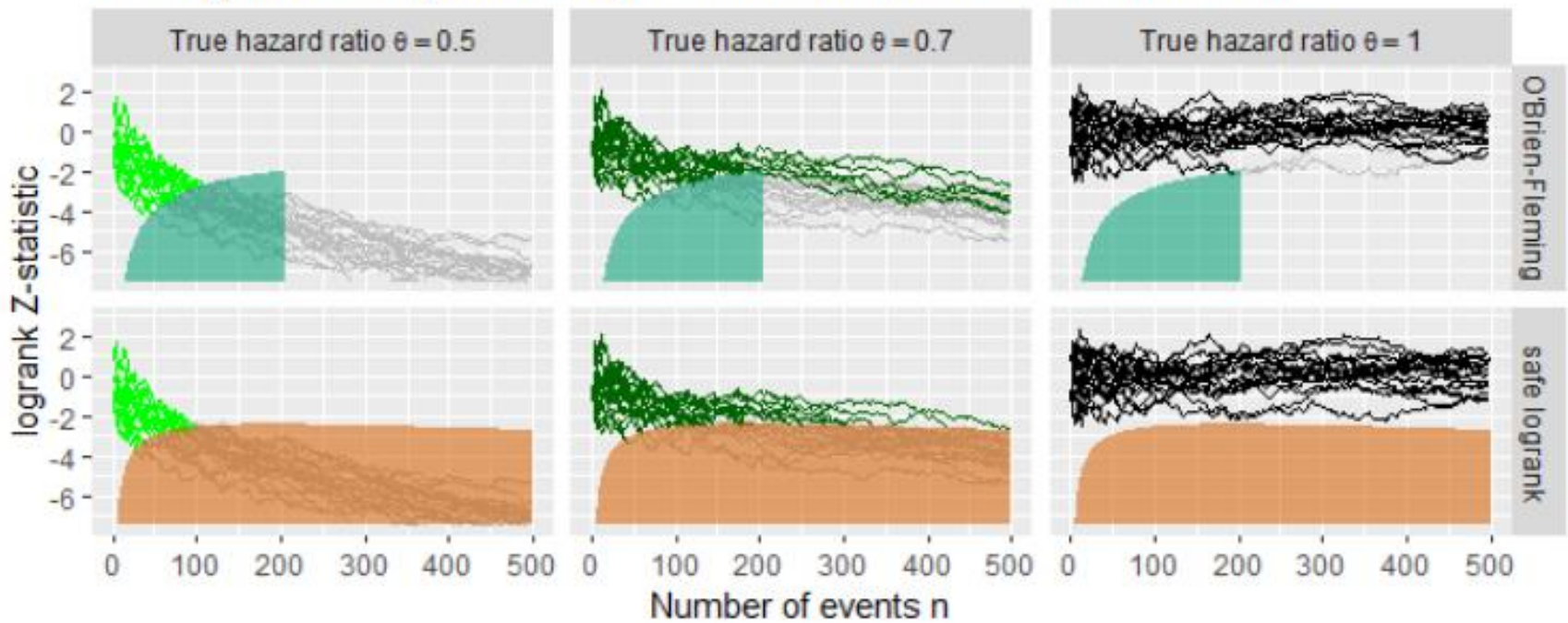
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Null hypothesis rejections by simulated data for a left-sided test



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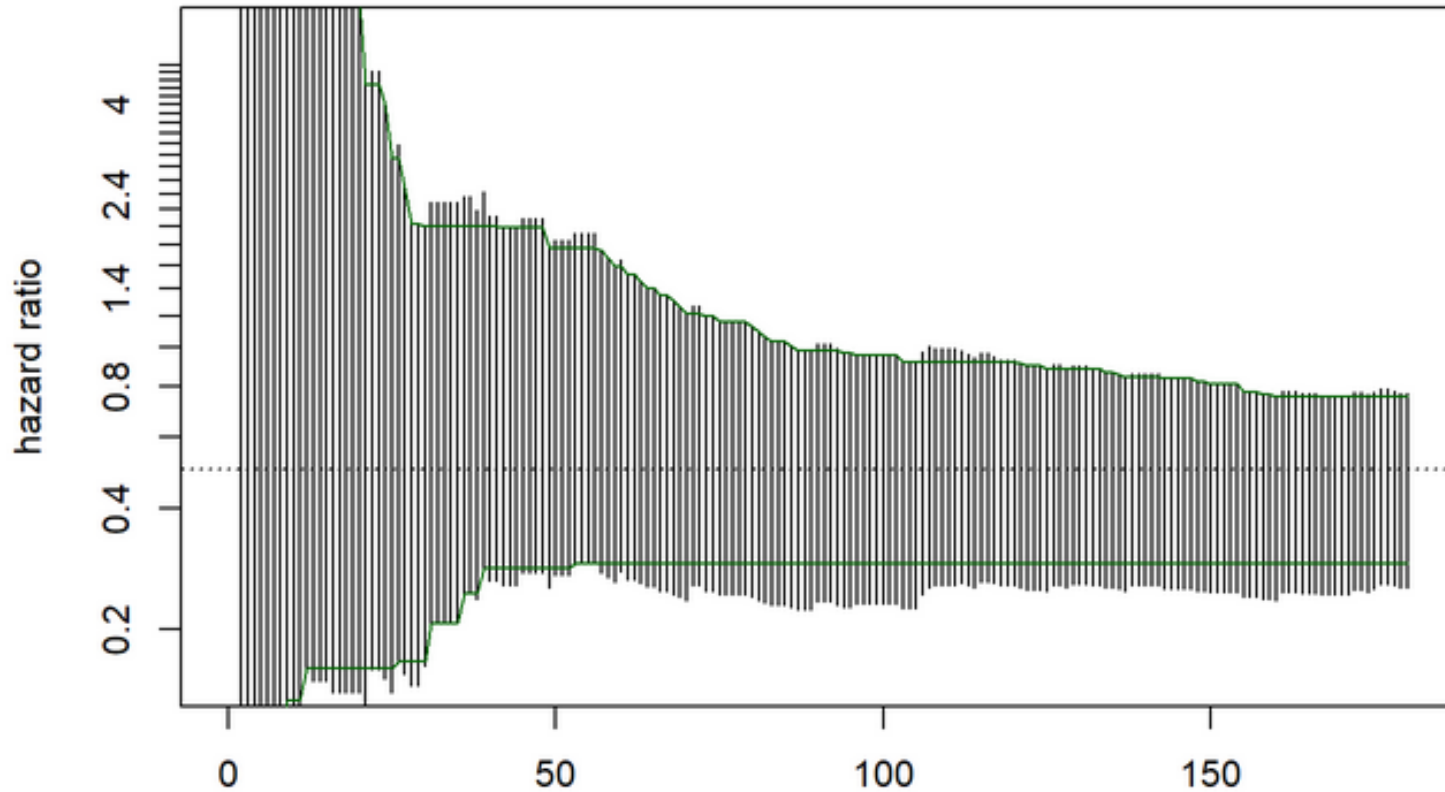
Null hypothesis rejections by simulated data for a left-sided test



H_0 not (yet) rejected: — $H_1: \theta = 0.5$ — $H_1: \theta = 0.7$ — $H_0: \theta = 1$ — rejected

Anytime-valid confidence sequences

Study 1: hazard ratio = 0.5



What are we assuming? Points for discussion

Heterogeneity

Simplicity

Anytime-valid confidence sequences

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- Future work confidence sequences based on random-effects model

References

- ter Schure, J.A., Pérez-Ortiz, M.F., Ly, A., & Grünwald, P.D. (2021). The Safe Logrank Test: Error Control under Continuous Monitoring with Unlimited Horizon. [arXiv:2011.06931](https://arxiv.org/abs/2011.06931)
- ALL-IN meta-analysis webinar (2020) <https://ir.cwi.nl/pub/30338>
originally for trials contributing to ALL-IN-META-BCG-CORONA
- ter Schure, J.A, & Grünwald, P.D. (2019). Accumulation Bias in meta-analysis: the need to consider time in error control [version 1; peer review: two approved]. *F1000Research*, 8. [doi:10.12688/f1000research.19375.1](https://doi.org/10.12688/f1000research.19375.1)
- Peto, Richard. Why do we need systematic overviews of randomized trials? ([Transcript of an oral presentation, modified by the editors](#)). *Statistics in medicine* 6.3 (1987): 233-240.

References

- ter Schure, J.A., Pérez-Ortiz, M.F., Ly, A., & Grünwald, P.D. (2021). The Safe Logrank Test: Error Control under Continuous Monitoring with Unlimited Horizon. [arXiv:2011.06931](https://arxiv.org/abs/2011.06931)
- ALL-IN meta-analysis webinar (2020) <https://ir.cwi.nl/pub/30338>
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